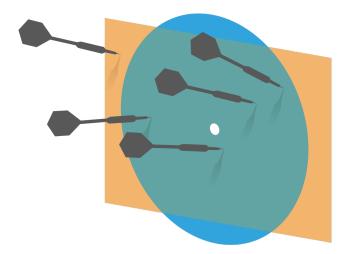
# Probability and Statistics for Computer Science



$$cov(X,Y) = E[(X - E[X])(Y - E[Y])]$$
$$= E[XY] - E[X]E[Y]$$

Covariance is coming back in matrix!

Credit: wikipedia

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### Last time

### Maximum likelihood Estimation (MLE II)

### # Bayesian Inference (MAP)

### Objective

### Review of Bayesian inference

# Wisualizing high dimensional data & Summarizing data

### \* The covariance matrix

### Refresh of some linear algebra

### Beta distribution

A distribution is Beta distribution if it has the following ▓ pdf:  $P(\theta) = K(\alpha, \beta)\theta^{\alpha-1}(1-\theta)^{\beta-1}$  $0 \le \Theta \le 1$ α >0, β>0 = 0 O.W. pdf of Beta – distribution 9 Beta(1,1)  $K(\alpha,\beta) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}$ Beta(5,5) Beta(50,50) Beta(70,70) Beta(20,50) ω Beta(0.5.0.5 Is an expressive family of ⋙ 9 density distributions 4  $\#Beta(\alpha = 1, \beta = 1)$  is uniform ΩI 0

0.0

0.2

0.4

0.6

θ

0.8

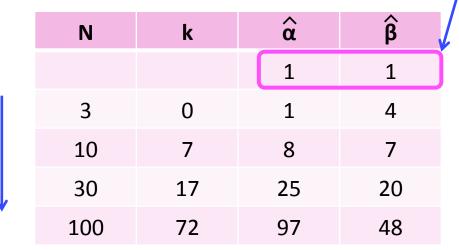
1.0

## Beta distribution as the conjugate prior for Binomial likelihood

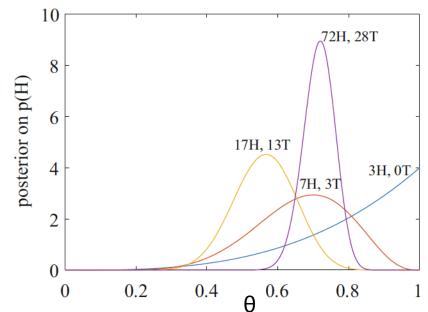
- \*\* The likelihood is Binomial (*N*, *k*)  $P(D|\theta) = \binom{N}{k} \theta^k (1-\theta)^{N-k}$
- \* The Beta distribution is used as the prior  $P(\theta) = K(\alpha,\beta)\theta^{\alpha-1}(1-\theta)^{\beta-1}$
- \* So  $P(\theta|D) \propto \theta^{\alpha+k-1}(1-\theta)^{\beta+N-k-1}$
- \* Then the posterior is  $Beta(\alpha + k, \beta + N k)$  $P(\theta|D) = K(\alpha + k, \beta + N - k)\theta^{\alpha + k - 1}(1 - \theta)^{\beta + N - k - 1}$

### The update of Bayesian posterior

- Since the posterior is in the same family as the conjugate prior, the posterior can be used as a new prior if more data is observed.
  - Suppose we start with a uniform prior on the probability  $\theta$  of heads



⊯



#### Maximize the Bayesian posterior (MAP)

\* The posterior of the previous example is

$$P(\theta|D) = K(\alpha + k, \beta + N - k)\theta^{\alpha + k - 1}(1 - \theta)^{\beta + N - k - 1}$$

Differentiating and setting to 0 gives the MAP estimate

$$\hat{\theta} = \frac{\alpha - 1 + k}{\alpha + \beta - 2 + N}$$

# Conjugate prior for other likelihood functions

- If the likelihood is Bernoulli or geometric, the conjugate prior is Beta
- If the likelihood is Poisson or Exponential, the conjugate prior is Gamma
- If the likelihood is normal with known variance, the conjugate prior is normal

#### A data set with high dimensions

### Seed data set from the UCI Machine Learning site:

	areaA	perimeterP	compactness	lengthKernel	widthKernel	asymmetry	lengthGroove	Label
1	15.26	14.84	0.871	5.763	3.312	2.221	5.22	1
2	14.88	14.57	0.8811	5.554	3.333	1.018	4.956	1
3	14.29	14.09	0.905	5.291	3.337	2.699	4.825	1
4	13.84	13.94	0.8955	5.324	3.379	2.259	4.805	1
5	16.14	14.99	0.9034	5.658	3.562	1.355	5.175	1
6	14.38	14.21	0.8951	5.386	3.312	2.462	4.956	1
7	14.69	14.49	0.8799	5.563	3.259	3.586	5.219	1

#### Matrix format of a dataset in the textbook

### Scatterplot matrix

- Wisualizing high dimensional data with scatter plot matrix
- Limited to  $\ast$ small number of scatter plots

Red: seed type I Blue: seed type II Yellow: seed type III 210 data points 7 dimensions

12 16 20

areaA

20 16

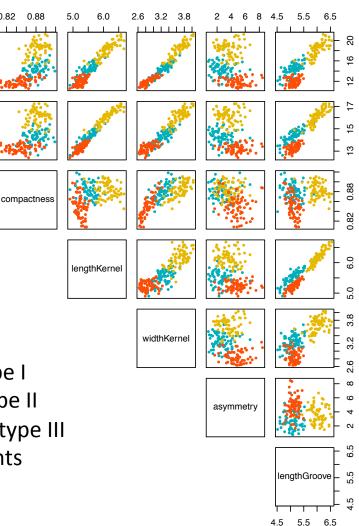
⊵

13

15 17

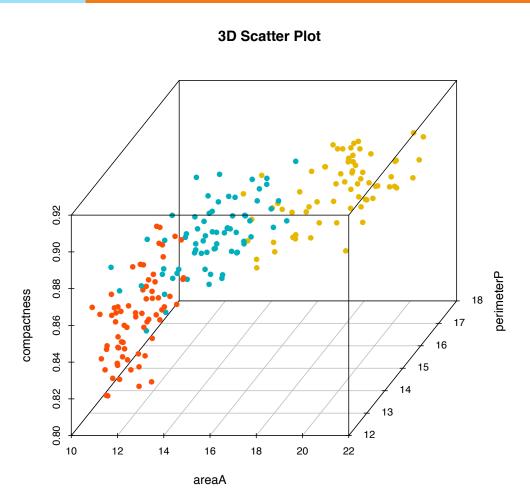
perimeterP

0.82



### 3D scatter plot

- We can also view
   the data set in 3
   dimensions
- But it's still
   limited in terms
   of number of
   dimensions we
   can see.



### Summarizing multidimensional data

- \* Location and spread parameters of a data set
- \* Notation
  - Write {x} for a dataset consisting of N data items
  - # Each item x<sub>i</sub> is a **d**-dimensional vector; column
  - **Write jth component of**  $x_i$  **as**  $x_i^{(j)}$ **; row**
  - Matrix for the data set {x} is d by N dimension

### Mean of a multidimensional data

We compute the mean of {x} by computing the mean of each component separately and stacking them to a vector

mean of jth component 
$$= \frac{\sum_i x_i^{(j)}}{N}$$

We write the mean of {x} as

$$mean(\{x\}) = \frac{\sum_i x_i}{N}$$

### Covariance

# \* The covariance of random variables X and Y is

### cov(X, Y) = E[(X - E[X])(Y - E[Y])]

### Note that

 $cov(X, X) = E[(X - E[X])^2] = var[X]$ 

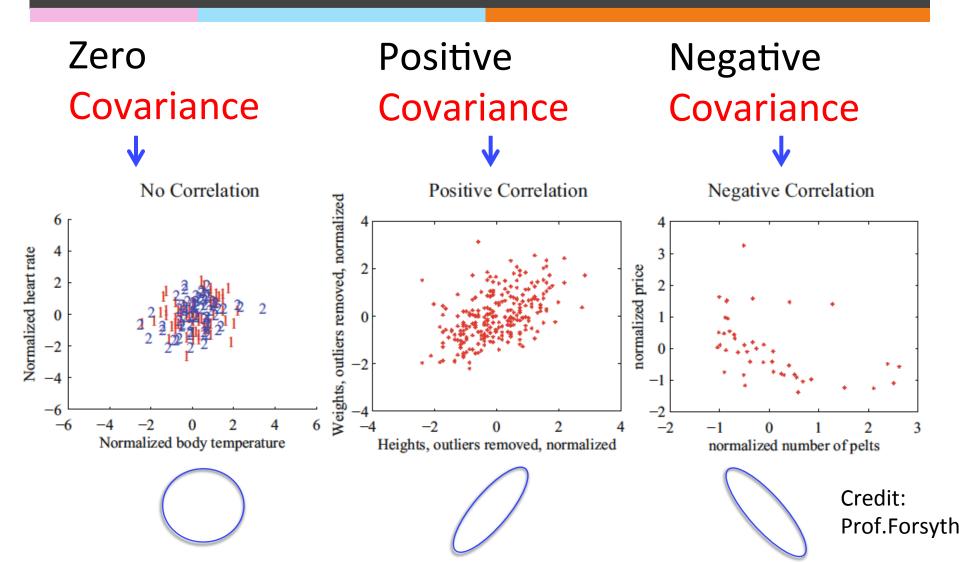
### Correlation coefficient is normalized covariance

\* The correlation coefficient is

$$corr(X,Y) = \frac{cov(X,Y)}{\sigma_X \sigma_Y}$$

When X, Y takes on values with equal probability to generate data sets {(x,y)}, the correlation coefficient will be as seen in Chapter 2.

### Covariance seen from scatter plots



### Covariance for a pair of components in a data set

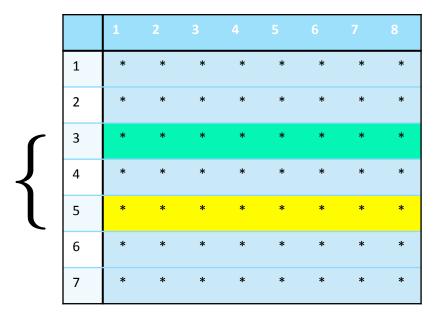
For the jth and kth components of a data set {x}

$$cov(\{x\}; j, k) = \frac{\sum_{i} (x_{i}^{(j)} - mean(\{x^{(j)}\}))(x_{i}^{(k)} - mean(\{x^{(k)}\}))^{T}}{N}$$

### Covariance of a pair of components

Data set 
$$ig\{\mathbf{X}ig\}$$
 7×8

 $cov({\mathbf{x}}; 3, 5)$ 

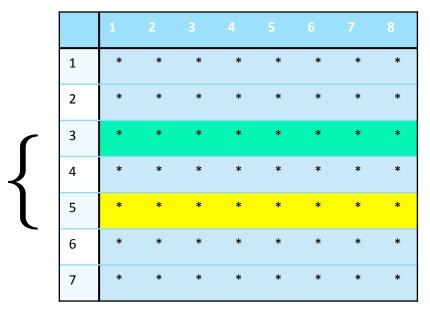


Take each row (component) of a pair and subtract it by the row mean, then do the inner product of the two resulting rows and divide by the number of columns

### Covariance of a pair of components

Data set 
$$\left\{ \mathbf{X} 
ight\}$$
 7×8

 $cov({\mathbf{x}}; 3, 5)$ 



How many pairs of rows are there for which we can compute the covariance?

49

64

56

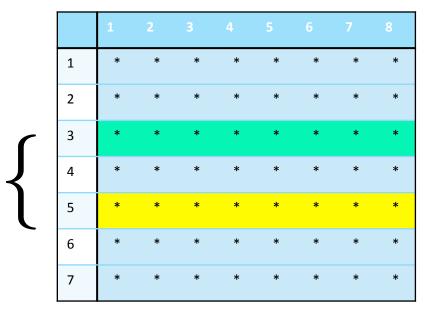
A)

B)

### Covariance matrix

Data set 
$$ig\{\mathbf{X}ig\}$$
 7×8

 $cov({\mathbf{x}}; 3, 5)$ 



Covmat(
$$\{\mathbf{X}\}$$
) 7×7

	1	2	3	4	5	6	7
1	*	*	*	*	*	*	*
2	*	*	*	*	*	*	*
3	*	*	*	*	*	*	*
4	*	*	*	*	*	*	*
5	*	*	*	*	*	*	*
6	*	*	*	*	*	*	*
7	*	*	*	*	*	*	*

### Properties of Covariance matrix

$$cov(\{x\}; j, j) = var(\{x^{(j)}\})$$
 Covmat( $\{\mathbf{x}\}$ ) 7×7

- The diagonal elements

   of the covariance matrix
   are just variances of
   each jth components
- The off diagonals are covariance between different components

	1	2	3	4	5	6	7
1	*	*	*	*	*	*	*
2	*	*	*	*	*	*	*
3	*	*	*	*	*	*	*
4	*	*	*	*	*	*	*
5	*	*	*	*	*	*	*
6	*	*	*	*	*	*	*
7	*	*	*	*	*	*	*

### Properties of Covariance matrix

$$cov(\{x\}; j, k) = cov(\{x\}; k, j)$$

Covmat(
$$\{\mathbf{X}\}$$
) 7×7

- \* The covariance matrix is symmetric!
- And it's positive semi-definite, that is all  $λ_i ≥ 0$
- Covariance matrix is diagonalizable

	1	2	3	4	5	6	7
1	*	*	*	*	*	*	*
2	*	*	*	*	*	*	*
3	*	*	*	*	*	*	*
4	*	*	*	*	*	*	*
5	*	*	*	*	*	*	*
6	*	*	*	*	*	*	*
7	*	*	*	*	*	*	*

### Properties of Covariance matrix

If we define x<sub>c</sub> as the mean centered matrix for dataset {x}

$$Covmat(\{x\}) = \frac{X_c X_c^T}{N}$$

\* The covariance matrix is a d×d matrix

Covmat(
$$\{\mathbf{x}\}$$
) 7×7

	1	2	3	4	5	6	7
1	*	*	*	*	*	*	*
2	*	*	*	*	*	*	*
3	*	*	*	*	*	*	*
4	*	*	*	*	*	*	*
5	*	*	*	*	*	*	*
6	*	*	*	*	*	*	*
7	*	*	*	*	*	*	*

$$A_0 = \begin{bmatrix} 5 & 4 & 3 & 2 & 1 \\ -1 & 1 & 0 & 1 & -1 \end{bmatrix} \begin{array}{c} \mathbf{X}^{(1)} \\ \mathbf{X}^{(2)} \end{array}$$

(I)

What are the dimensions of the covariance matrix of this data?

(I)  

$$A_{0} = \begin{bmatrix} 5 & 4 & 3 & 2 & 1 \\ -1 & 1 & 0 & 1 & -1 \end{bmatrix}$$

$$A_{1} = \begin{bmatrix} 2 & 1 & 0 & -1 & -2 \\ -1 & 1 & 0 & 1 & -1 \end{bmatrix}$$

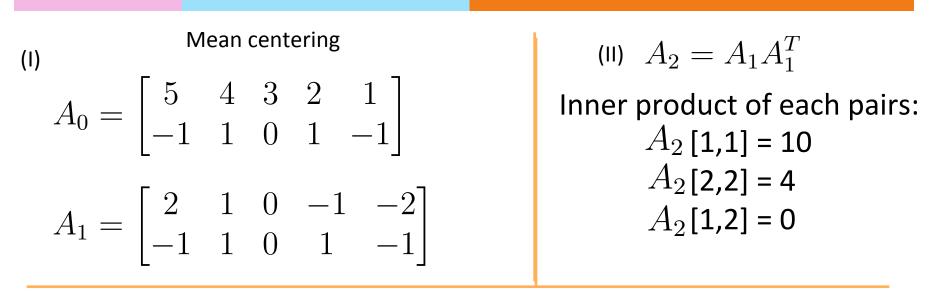
Mean centering  

$$A_{0} = \begin{bmatrix} 5 & 4 & 3 & 2 & 1 \\ -1 & 1 & 0 & 1 & -1 \end{bmatrix}$$

$$A_{1} = \begin{bmatrix} 2 & 1 & 0 & -1 & -2 \\ -1 & 1 & 0 & 1 & -1 \end{bmatrix}$$

(II) 
$$A_2 = A_1 A_1^T$$

Inner product of each pairs:  $A_2$  [1,1] = 10  $A_2$  [2,2] = 4  $A_2$  [1,2] = 0



#### (111)

Divide the matrix with N – the number of items

**Covmat({x})** = 
$$\frac{1}{N}A_2 = \frac{1}{5}\begin{bmatrix} 10 & 0\\ 0 & 4 \end{bmatrix} = \begin{bmatrix} 2 & 0\\ 0 & 0.8 \end{bmatrix}$$

### What do the data look like when Covmat({x}) is diagonal?

X<sup>(2)</sup>  $A_0 = \begin{bmatrix} 5 & 4 & 3 & 2 & 1 \\ -1 & 1 & 0 & 1 & -1 \end{bmatrix}$  $X^{(1)}$ \* Covmat({x}) =  $\frac{1}{N}A_2 = \frac{1}{5}\begin{bmatrix} 10 & 0\\ 0 & 4 \end{bmatrix} = \begin{bmatrix} 2 & 0\\ 0 & 0.8 \end{bmatrix}$ 

# Translation properties of mean and covariance matrix

\* Translating the data set translates the mean

$$mean(\{x\} + c) = mean(\{x\}) + c$$

\* Translating the data set leaves the covariance matrix unchanged

 $Covmat(\{x\}+c) = Covmat(\{x\})$ 

# Translation properties of covariance matrix



### Linear transformation properties of mean and covariance matrix

\* Linearly transforming the data set linearly transforms the mean

$$mean(\{A\mathbf{x}\}) = A mean(\{\mathbf{x}\})$$

Linearly transforming the data set linearly changes the covariance matrix quadratically

 $Covmat({Ax}) = A \ Covmat({x})A^T$ 

### Proof of linear transformation of covariance matrix

### **Dimension Reduction**

- In stead of showing more dimensions through visualization, it's a good idea to do dimension reduction in order to see the major features of the data set.
- \* For example, principal component analysis help find the major components of the data set.
- \* PCA is essentially about finding eigenvectors of covariance matrix

### Refresh of some linear algebra

### Why linear algebra?

- We are now into part IV of the course. The contents will be basic machine learning techniques.
- \* Linear algebra is essential for a lot of machine Learning methods!

#### Eigenvalues and eigenvectors review

- \* If A is an **n×n** square matrix, an eigenvalue  $\lambda$  and its corresponding eigenvector v (of dimension n×1) satisfy  $Av = \lambda v$ .
- \* To solve for  $\lambda$ , we solve the characteristic equation

$$|A - \lambda I| = 0$$

\* Given a value of  $\lambda$ , we solve v by solving

$$(A - \lambda I) v = 0$$

\* Note if v is an eigenvector, then so is any multiple kv.

Find the eigenvalues and eigenvectors

$$A = \begin{bmatrix} 5 & 3 \\ 3 & 5 \end{bmatrix}$$

Find the eigenvalues and eigenvectors

$$A = \begin{bmatrix} 5 & 3 \\ 3 & 5 \end{bmatrix}$$

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$$A = \begin{bmatrix} 5 & 3 \\ 3 & 5 \end{bmatrix}$$

$$A = \begin{bmatrix} 5 - \lambda & 3 \\ 3 & 5 - \lambda \end{bmatrix} = (5 - \lambda)^{2} - 3^{2} = \lambda^{2} - 10\lambda + 15 - 9$$

$$= \lambda^{2} - 10\lambda + 16 = 0$$

$$= \lambda^{2} - 10\lambda + 16 = 0$$

$$= (\lambda - 8)(\lambda - 2) = 0$$
So the eigenvertues  $\lambda_{1} = 8$ ,  $pos^{2} + ive = definite$ 

$$\lambda_{2} = 2$$
,  $pos^{2} + ive = definite$ 

Find the eigenvectors

$$A = \begin{bmatrix} 5 & 3\\ 3 & 5 \end{bmatrix}$$

% Find the
eigenvectors

$$A = \begin{bmatrix} 5 & 3 \\ 3 & 5 \end{bmatrix} \quad \begin{bmatrix} -\sigma & \lambda_1 = \delta & A - \delta 1 = \begin{pmatrix} 5 - \delta & 3 \\ 3 & 5 - \delta \end{pmatrix} = \begin{pmatrix} -3 & 3 \\ 7 & -3 \end{pmatrix}$$
$$(A - \delta 1) \cup_i = o$$
$$\Rightarrow \cup_i = \begin{bmatrix} i \\ i \end{bmatrix}$$
$$\begin{bmatrix} -\sigma & \lambda_2 = 2 & A - 21 = \begin{bmatrix} 5 - 2 & 3 \\ 3 & 5 - 2 \end{bmatrix} = \begin{bmatrix} 3 & 3 \\ 3 & 3 \end{bmatrix}$$
$$(A - 21) \cup_i = o$$
$$\Rightarrow \bigcup_i 2 = \begin{bmatrix} -1 \\ -1 \end{bmatrix}$$

Find the eigenvalues and eigenvectors of

$$A = \begin{bmatrix} 1 & 2\\ 2 & 4 \end{bmatrix}$$

Find the eigenvalues and eigenvectors of

$$A = \begin{bmatrix} 1 & 2\\ 2 & 4 \end{bmatrix}$$

$$A = \begin{bmatrix} 1 & 2 \\ 2 & 4 \end{bmatrix} A \text{ is symmetric}$$

$$|A - \lambda 1| = \begin{bmatrix} 1 - \lambda & 2 \\ 2 & 4 - \lambda \end{bmatrix} = (1 - \lambda) (4 - \lambda) - 4$$

$$= \lambda^{2} - 5\lambda = 0$$
So che eigenvalues are  $\lambda_{1} = 5, \lambda_{2} = 0$ 

$$\int Politive \quad \text{femi-definite}$$

\* Find the eigenvectors of  $A = \begin{bmatrix} 1 & 2 \\ 2 & 4 \end{bmatrix}$ 

\* Find the eigenvectors of  $A = \begin{bmatrix} 1 & 2 \\ 2 & 4 \end{bmatrix}$ 

For 
$$\lambda_{1}=5$$
  $A-51 = \begin{bmatrix} -4 & 2 \\ 2 & -1 \end{bmatrix}$   
 $(A-51) V_{1} = 0$   
 $\Rightarrow V_{1} = \begin{bmatrix} 2 \\ 2 \end{bmatrix} \Rightarrow U_{1} = \int_{-5}^{1} \begin{bmatrix} 2 \\ 2 \end{bmatrix}$   
 $\lambda_{2}=0$   $A V_{2} = 0$   
 $\Rightarrow \begin{bmatrix} 1 & 2 \\ 2 & 4 \end{bmatrix} V_{2} = 0$   
 $\Rightarrow V_{2} = \begin{bmatrix} -2 \\ 1 \end{bmatrix} \Rightarrow U_{2} = \int_{-2}^{1} \begin{bmatrix} -2 \\ -2 \end{bmatrix}$ 

#### Diagonalization of a symmetric matrix

- If A is an n×n symmetric square matrix, the eigenvalues are real.
- If the eigenvalues are also distinct, their eigenvectors are orthogonal
- \* We can then scale the eigenvectors to unit length, and place them into an orthogonal matrix  $U = [\mathbf{u}_1 \, \mathbf{u}_2 \, ..., \, \mathbf{u}_n]$
- \* We can write the diagonal matrix  $\Lambda = U^T A U$  such that the diagonal entries of  $\Lambda$  are  $\lambda_1, \lambda_2 \dots \lambda_n$  in that order.

## Diagonalization example

**⊮** For

$$A = \begin{bmatrix} 5 & 3 \\ 3 & 5 \end{bmatrix}$$

For 
$$\lambda_1 = 8$$
  $A - 82 = \begin{pmatrix} 5 - 8 & 3 \\ 3 & 5 - 8 \end{pmatrix} = \begin{pmatrix} -3 & 3 \\ 3 & -3 \end{pmatrix}$   
 $(A - 81) \cup_1 = 0$   
 $\Rightarrow \cup_1 = \begin{bmatrix} 1 \\ 1 \end{bmatrix}$   
For  $\lambda_2 = 2$   $A - 21 = \begin{bmatrix} 5 - 2 & 3 \\ 3 & 5 - 2 \end{bmatrix} = \begin{bmatrix} 3 & 3 \\ 3 & 3 \end{bmatrix}$   
 $(A - 21) \cup_2 = 0$   
 $\Rightarrow \bigcup_2 = \begin{bmatrix} 1 \\ -1 \end{bmatrix}$ 

$$\begin{split} \lambda_{i} = \delta_{i} \Rightarrow \nu_{i} = \begin{bmatrix} i \\ j \end{bmatrix} \Rightarrow u_{i} = \frac{1}{\|\nu_{i}\|} \nu_{i} = \frac{1}{2^{2}} \begin{bmatrix} i \\ j \end{bmatrix} \\ \lambda_{2} = 2 \Rightarrow \nu_{2} = \begin{bmatrix} i \\ -i \end{bmatrix} \Rightarrow u_{2} = \frac{1}{\|\nu_{2}\|} \nu_{2} = \frac{1}{2^{2}} \begin{bmatrix} i \\ -i \end{bmatrix} \\ \begin{bmatrix} \delta \\ 0 \end{bmatrix} = \begin{bmatrix} \frac{1}{2^{2}} & \frac{1}{2^{2}} \end{bmatrix} \begin{bmatrix} 5 & 3 \\ \frac{1}{2^{2}} & -\frac{1}{2^{2}} \end{bmatrix} \begin{bmatrix} \frac{1}{2^{2}} & \frac{1}{2^{2}} \\ \frac{1}{2^{2}} & -\frac{1}{2^{2}} \end{bmatrix} \begin{bmatrix} \frac{1}{2^{2}} & \frac{1}{2^{2}} \\ \frac{1}{2^{2}} & -\frac{1}{2^{2}} \end{bmatrix} \end{split}$$

#### Q. Are these two vectors orthogonal?

#### Q. Is this true?

# When two zero-mean vectors of data are orthogonal, they are uncorrelated

A. Yes

B. No

## See you next time

See You!

